Multivariate Statistical Analysis

Lecture 10

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Outline

Sample Correlation Coefficient

2 Tests for the Hypothesis of Lack of Correlation

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Given the sample $\mathbf{x}_1, \dots, \mathbf{x}_N$ from $\mathcal{N}_p(\mu, \mathbf{\Sigma})$, the maximum likelihood estimator of the correlation between the *i*-th and the *j*-th components is

$$r_{ij} = \frac{\sum_{\alpha=1}^{N} (x_{i\alpha} - \bar{x}_i)(x_{j\alpha} - \bar{x}_j)}{\sqrt{\sum_{\alpha=1}^{N} (x_{i\alpha} - \bar{x}_i)^2} \sqrt{\sum_{\alpha=1}^{N} (x_{j\alpha} - \bar{x}_j)^2}},$$

where $x_{i\alpha}$ is the *i*-th component of \mathbf{x}_{α} and

$$\bar{x}_i = \frac{1}{N} \sum_{\alpha=1}^N x_{i\alpha}.$$

We shall find the distribution of r_{ij} .

If the population correlation

$$\rho_{ij} = \frac{\sigma_{ij}}{\sqrt{\sigma_{ii}\sigma_{jj}}}$$

is zero, then the density of sample correlation r_{ij} is

$$k_N(r) = \frac{\Gamma(\frac{N-1}{2})}{\sqrt{\pi} \Gamma(\frac{N-2}{2})} (1 - r_{ij}^2)^{\frac{N-4}{2}}.$$

Let $\mathbf{x}_1, \dots, \mathbf{x}_N$ be observation from $\mathcal{N}_2(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, where

$$m{\mu} = egin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix}$$
 and $m{\Sigma} = egin{bmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{12} & \sigma_{22} \end{bmatrix}$

We denote

$$\mathbf{x}_{\alpha} = \begin{bmatrix} x_{1\alpha} \\ x_{2\alpha} \end{bmatrix}, \quad \bar{\mathbf{x}} = \frac{1}{N} \sum_{\alpha=1}^{N} \mathbf{x}_{\alpha} \quad \text{and} \quad \mathbf{A} = \sum_{\alpha=1}^{N} (\mathbf{x}_{\alpha} - \bar{\mathbf{x}})(\mathbf{x}_{\alpha} - \bar{\mathbf{x}})^{\top}.$$

We have shown that A can be written as

$$\mathbf{A} = egin{bmatrix} a_{11} & a_{12} \ a_{21} & a_{22} \end{bmatrix} = \sum_{lpha=1}^n \mathbf{z}_lpha \mathbf{z}_lpha^ op,$$

where n = N-1 and $\mathbf{z}_1, \dots, \mathbf{z}_n$ are independent distributed to $\mathcal{N}_2\left(\mathbf{0}, \mathbf{\Sigma}\right)$

We denote

$$a_{11.2} = a_{11} - \frac{a_{12}^2}{a_{22}}, \qquad \sigma_{11.2} = \sigma_{11} - \frac{\sigma_{12}^2}{\sigma_{22}} \quad \text{and} \quad r = \frac{a_{12}}{\sqrt{a_{11}}\sqrt{a_{22}}}.$$

Lemma

Based on above notations, we have

(a)
$$\frac{a_{11}}{\sigma_{11}} \sim \chi_n^2$$
 and $\frac{a_{22}}{\sigma_{22}} \sim \chi_n^2$;

(b)
$$a_{12} \mid a_{22} \sim \mathcal{N}\left(\sigma_{12}\sigma_{22}^{-1}a_{22}, \sigma_{11.2}a_{22}\right)$$
;

(c)
$$\frac{a_{11.2}}{\sigma_{11.2}} \sim \chi^2_{n-1}$$
 is independent on a_{12} and a_{22} .

We can show that

$$z = \frac{x}{\sqrt{y/(n-1)}}$$
$$= \frac{\sqrt{n-1}(r - \sigma_{12}\sigma_{22}^{-1}\sqrt{a_{22}/a_{11}})}{\sqrt{1-r^2}}$$

where

$$x = \frac{a_{12} - \sigma_{12}\sigma_{22}^{-1}a_{22}}{\sqrt{\sigma_{11.2}a_{22}}} \sim \mathcal{N}(0,1)$$
 and $y = \frac{a_{11.2}}{\sigma_{11.2}} \sim \chi_{n-1}^2$

are independent.

If
$$\sigma_{12}=0$$
, then $z=rac{x}{\sqrt{y/(n-1)}}\sim t_{n-1}$.

If population correlation

$$\rho = \frac{\sigma_{12}}{\sqrt{\sigma_{11}\sigma_{22}}}$$

is non-zero ($\sigma_{12} \neq 0$), the density of sample correlation r is

$$\frac{2^{n-2}(1-\rho^2)^{\frac{n}{2}}(1-r^2)^{\frac{n-3}{2}}}{(n-2)!\pi}\sum_{\alpha=0}^{\infty}\frac{(2\rho r)^{\alpha}}{\alpha!}\left(\Gamma\left(\frac{n+\alpha}{2}\right)\right)^2.$$

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Tests for the Hypothesis of Lack of Correlation

Consider the hypothesis $H: \rho_{ij} = 0$ for some particular pair (i,j).

• For testing H against alternatives $\rho_{ij} > 0$, we reject H if $r_{ij} > r_0$ for some positive r_0 . The probability of rejecting H when H is true is

$$\int_{r_0}^1 k_N(r) \, \mathrm{d}r.$$

- ② For testing H against alternatives $\rho_{ij} < 0$, we reject H if $r_{ij} < -r_0$.
- **3** For testing H against alternatives $\rho_{ij} \neq 0$, we reject H if $r_{ij} > r_1$ or $r_{ij} < -r_1$ for some positive r_1 . The probability of rejection when H is true is

$$\int_{-1}^{-r_1} k_N(r) \, \mathrm{d}r + \int_{r_1}^1 k_N(r) \, \mathrm{d}r.$$

Tests for Lack of Correlation

We have shown that

$$\sqrt{N-2} \cdot \frac{r_{ij}}{\sqrt{1-r_{ij}^2}}$$

has the *t*-distribution with N-2 degrees of freedom.

We can also use *t*-tables. For $\rho_{ij} \neq 0$, reject *H* if

$$\sqrt{N-2}\cdot\frac{|r_{ij}|}{\sqrt{1-r_{ij}^2}}>t_{N-2}(\alpha),$$

where $t_{N-2}(\alpha)$ is the two-tailed significance point of the t-statistic with N-2 degrees of freedom for significance level α .